Test-retest reliability of the Test of Infant Motor Performance Screening Items in infants at risk for impaired functional motor performance

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Abstract

Objective: To examine test-retest reliability of the TIMPSI in infants at risk for impaired functional motor performance.  
Methods: The TIMPSI was administered twice to 51 infants from two different hospitals in Norway. 
Results: The intra-class correlation coefficient was 0.99. 
Conclusion: Test-retest reliability of the TIMPSI was excellent.

Keywords

Test-retest, preterm / high-risk infants, motor assessment.

Conflict of interest

Dr. Campbell and Dr. Girolami are co-developers of the TIMPSI and partners in Infant Motor Performance Scales, LLC, the publisher of the TIMPSI.
1. Introduction

Motor assessments in infants at risk for developmental delay are primarily performed to discriminate between typically developing infants and infants with suspected neurological dysfunction. This is important when planning intervention, predicting motor difficulties and evaluating change over time [1, 2]. In order to direct resources towards infants likely to gain most from intervention, while avoiding intervention on infants with typical development, it is essential that assessment tools are reliable and valid. The prevalence of developmental difficulties in infants born preterm increases with decreasing gestational age at birth (GA) [3], and the incidence of motor disabilities such as cerebral palsy and developmental coordination disorder is particularly high [3-5]. Systematic reviews of neonatal assessments tools conclude that the Test of Infant Motor Performance (TIMP) is one of the best motor assessment tools to discriminate between infants with age-appropriate motor development and infants with delayed motor performance. Further it is also useful for planning interventions and evaluating change over time [1, 2, 6, 7].

The TIMP was developed to assess functional motor performance in new-borns and infants from 34 weeks postmenstrual age (PMA) to 17 weeks corrected age (CA). Conducted at 3 months CA the TIMP was predictive of children’s motor performance at 4-5 years, as measured by The Peabody Developmental Motor Scales [8]. A test-retest reliability study of the TIMP in 106 infants PMA 32 weeks to CA 16 weeks with varying risk and ethnicity demonstrated a high correlation between scores on two different days ($r = 0.89$) [7].

Average time to conduct the TIMP is 25-35 minutes, which for the youngest and most fragile infants may be too demanding. Therefore, a shorter version was developed, the Test of Infant Motor Performance Screening Items (TIMPSI) [6, 9], for identifying infants for whom the full version should be conducted. In a group of low birth weight infants between PMA 34 weeks and CA 17 weeks total TIMPSI scores correlated well with total TIMP scores ($r = 0.88$) [9]. A test-retest reliability study of the TIMPSI in infants at risk for long-term motor difficulties has not yet been carried out, but should be performed before routinely implementing this test for the assessment of fragile infants. We wanted to explore the clinical utility of the TIMPSI by investigating the stability in scores and the measurement error when the same tester conducted two consecutive tests. The aim of this study was to examine test-retest reliability of the TIMPSI in a group of infants at high to moderate risk for long-term motor developmental difficulties.
2. Methods and participants

This study used an observational design to investigate test-retest reliability of the TIMPSI within a period of three days. This specific time frame was chosen because developmental changes are expected to be minimal over such a short interval [7]. In order to generalise the findings to clinical work, infants between PMA 36-37 weeks and CA 12-13 weeks, with varying risk for neurologic diagnosis or motor delay, were recruited.

Between April 2013 to December 2014, fifty-one infants from two hospitals in Norway, the University Hospital of North Norway (n=14) and St.Olavs Hospital, Trondheim University Hospital (n=37), were recruited for this study. Infants with high or moderate risk for long term motor development difficulties were eligible for inclusion. High risk was defined as infants born prior to 28 weeks GA with a birth weight <1000 grams, infants with Grade III or IV intraventricular haemorrhages or periventricular leukomalacias and term infants with severe asphyxia treated with hypothermia. Moderate risk was defined as GA from 28 to 33 weeks. Parents were required to understand Norwegian or English. Medically unstable infants, infants who had undergone surgery and infants with genetic syndromes were excluded. With the exception of holidays and periods when the testers were on leave, eligible infants were continuously recruited. The sample was a convenience sample depending on availability of infants and parents at two time points as well as testers.

The study protocol was reviewed by the Regional Committees for Medical and Health Research Ethics (REC) January 2012, which concluded that the study did not require approval but should be reported to the Data Protection Officer at the Hospital.

2.2. The assessment tool

The TIMPSI is comprised of 29 of the 42 items from the TIMP. There are observed items scored during the observation of spontaneous movements and elicited items designed to assess the responses to visual and auditory stimuli, handling and changes of position [6]. The test is divided into three subsets: a Screening set, an Easy set, and a Hard set. The Screening set consists of 11 items with rating scales from five- to seven-points, score range 0-51. All infants are first assessed with the Screening Set. Based on the raw score of the Screening Set, a second set of either 10 easier or 8 harder items is administered to obtain a total score for motor performance [6]. The Easy set has four dichotomously scored items and six items with a five- or six-point rating scale, score range 0-31. The
Hard set has eight items: five dichotomously scored and three with a five-point rating scale, score range 0-17 [10]. The scores for the administered items are summed with higher scores indicative of better motor performance, maximum score 99. TIMPSI age standards are available in the TIMP manual [6] based on the motor performance of 990 U.S. infants [9]. Average scores for infants PMA 36-37 weeks is 42 (SD: 16) and 79 (SD: 13) for infants CA 12-13 weeks.

2.3. Procedure

One tester from each hospital participated. Both testers were experienced paediatric physiotherapists who had attended workshops on the TIMP and had been using the test regularly for several years. A physiotherapist unknown to the parents in the NICU or Follow-up clinic invited all parents of eligible infants to participate in the study and a written consent was obtained. Because we aimed to minimize the burden for each infant and parents, test 1 was administered as part of ordinary clinical practice, either at week 36–37 PMA or at week 12-13 CA. Approximately half of the infants were tested at week 36-37 PMA and half tested at week 12-13 CA. The infants should be in “State of arousal level “ three (eyes open, no movements) or four (eyes open, large movements) according to Prechtl ‘s States [11]. The ideal time of the day for most of the infants was following a period of sleep and before meals. Test 2 was carried out within three days after test 1. In case of two tests carried out on the same day, pauses of several hours between the tests ensured the infants were rested and in the proper behavioural state for testing. In addition, testers would not remember scoring details of the previous test.

2.4. Statistical analysis

Sample size was estimated a priori according to Walter [12]. With a power of 80% and a significance level of 5%, we needed 45 participants to achieve an intra-class correlation coefficient (ICC) ≥0.8. Normality of the data was examined by the Shapiro-Wilk test. Relative reliability between Test 1 and Test 2 for within-subject differences was assessed by calculating ICC1,1 [13]. Relative reliability refers to consistent ranking of scores for an individual in a group by repeated measurements. Absolute reliability, the standard error of measurement, was calculated as the square root of the mean within-subject variance (SW) [14, 15]. SW is expressed in the original measurement scale with a low value expressing a small degree of measurement error. The difference between a subject’s measurement and the true value would be expected to be less than 1.96 x SW for 95% of the observations [14]. The
difference between the two measurements for the same subject is then expected to be less than \( \sqrt{2} \times 1.96 \times S_W = 2.77 \times S_W \) for 95% of the pairs of observations [14]. Bland Altman plot was used for verifying the consistency of the measurements [16]. This plot gives a graphical presentation of the differences between two tests plotted against the mean difference of the two tests allowing visual assessment of the scoring distribution and potential measurement bias [16]. The software IBM SPSS statistics version 22 was used to perform the statistical analyses.

3. Results

The mean time interval between Test 1 and Test 2 was 1 day (SD: 0.84). Thirteen (25%) of the infants had both tests administrated the same day. Sample characteristics are presented in Table 1. Forty-five (88%) of the infants were born at or before week 33 GA and six (12%) were born at term. Thirty-two (63%) were boys. One in four had abnormal caput ultrasound, most of them in the high risk group, and one in four had bronchopulmonary dysplasia.

The infants tested at week 36-37 PMA scored in the below average range based on the normative sample with an average total raw score of 31.8 (SD 10.3) [6]. Infants tested at week 12-13 CA scored within the average range for age, with an average total score of 76.1 (SD 7.4). Since in general the scores were normally distributed, ICC_{1,1} could be used to assess the degree of correlation between repeated tests. ICC_{1,1} was 0.99 for all infants, 0.94 for infants tested at week 36-37 PMA, and 0.93 for infants tested at week 12-13 CA (Table 2).

Measurement error (\( S_W \)) for the total TIMPSI score of all infants was 3.1 which implies that in 95% of the cases the measurement error will be within 3.1 \( \times 1.96 \) which equals 6.07 points on the total TIMPSI score. When comparing two time points, the difference should exceed the smallest detectable difference (SDD) calculated as \( 2.77 \times S_W \) to be sure that there is a difference beyond measurement error [14]. For our study this means that the differences in scores would need to exceed 9.7 points in infants with PMA 36-37 weeks and 7.1 points in infants with CA 12-13 weeks to indicate that real change has occurred.

The Bland Altman plot (Figure 1) shows the agreement between the tests at two time points. The mean differences of the two tests was close to zero which indicates a very high agreement. The scores from 48 (94%) infants fell within 1.96 standard deviations of the mean difference for all observations equally distributed above and below the zero point. Upon visual inspection no differences
were found between infants tested on the same day and infants tested on two different days.

4. Discussion

This is the first test-retest reliability study of the Test of Infant Motor Performance Screening Items (TIMPSI) in infants at risk for developmental problems. A previous test-retest reliability study of the full version of the TIMP demonstrated high correlation using Pearson’s $r$ (0.89) [7]. A test-retest study of the TIMPSI in a study population of children with spinal muscular atrophy also demonstrated high correlation using Pearson’s $r$ (0.95) [10]. Pearson’s $r$ is a measure of linear correlation between two values [17], while the ICC provides estimates of both association and agreement. Because we used the ICC$_{1,1}$ we cannot directly compare our results with the two aforementioned studies. Our results showed ICC values ≥ 0.93, which indicate excellent relative reliability of the TIMPSI. Values of 0.7 to 0.8 are regarded as satisfactory but for clinical application values of 0.90 are desirable [18].

Additionally, we calculated absolute reliabilities by $S_W$, which was high, implying that for evaluative purposes change in total scores must be rather high to conclude that there have been real changes beyond typical development. Our results are consistent with the purpose of the TIMPSI, which is to screen development in order to determine whether a full TIMP should be administered for discriminative purposes. Furthermore, the TIMP rather than the TIMPSI should be used to evaluate changes over time.

The spread of the scores in the Bland Altman plot was evenly distributed with approximately 95% within the limit of agreement. Two of the three infants that fell outside the limits, were tested at PMA 36-37 weeks and had low scores on the two tests. This might indicate that for subjects with low scores there is less consistency. However, due to the low number of infants this finding cannot be generalized.

Agreements and associations using the ICC in a group of infants with high to moderate risk of adverse neurodevelopment have not previously been reported. High ICC is likely if the infants’ behavioural state is the same and the testers have high intra-rater reliability. Based on the ICC in this study, correspondence between the two tests was excellent, but the measurement errors were high. Variability between two tests can be caused by the instrument, the tester or the subject being tested. The infants were tested when they were in a satisfactory behavioural state and none or minimal change in motor development was expected during this period between tests.
One limitation of this study may be that some infants were assessed twice during one day. This is not ideal, but for participants living far from the hospital, this was the only possibility. When performing two tests on the same day, care was taken not to know the results from the first test when conducting the second test, for example, by assessing other infants between the two tests. The visual inspection of the data showed no differences between infants tested on the same or separate days.

Both testers were experienced paediatric physiotherapists with thorough knowledge of the test. Because only one tester from each hospital participated and the two hospitals are not located in the same area of the country, the testers assessed different babies. Consequently, we were unable to assess inter-rater reliability, which would have strengthened this study.

5. Conclusion

The TIMPSI showed strong test-retest reliability when performed on a group of infants with high to moderate risk for later motor developmental difficulties. We can recommend use of the TIMPSI to screen development of infants for whom the full version of the test is too demanding.

Acknowledgment

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References

Table 1. Neonatal characteristics and age of subjects tested using the TIMPSI.

<table>
<thead>
<tr>
<th></th>
<th>High risk (n=27)</th>
<th>Moderate risk (n=24)</th>
<th>Total (n=51)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Birth weight (grams): mean (SD)</td>
<td>1499 (1158)</td>
<td>1546 (292)</td>
<td>1524 (814)</td>
</tr>
<tr>
<td>Gestational age at birth (weeks): mean (SD)</td>
<td>29.8 (6.2)</td>
<td>30.4 (1.7)</td>
<td>30.1 (4.4)</td>
</tr>
<tr>
<td>Bronchopulmonary dysplasia: n (%)</td>
<td>12 (24%)</td>
<td>0 (0%)</td>
<td>12 (24%)</td>
</tr>
<tr>
<td>Abnormal caput ultrasound: n (%)</td>
<td>9 (18%)</td>
<td>4 (8%)</td>
<td>13 (25%)</td>
</tr>
<tr>
<td>Intracranial bleed Grade III or IV: n (%)</td>
<td>2 (4%)</td>
<td>0 (0%)</td>
<td>2 (4%)</td>
</tr>
<tr>
<td>Periventricular leucomalasia: n (%)</td>
<td>3 (6%)</td>
<td>2 (4%)</td>
<td>5 (10%)</td>
</tr>
<tr>
<td>Infants tested at postmenstrual age 36-37 weeks: n (%)</td>
<td>6 (12%)</td>
<td>21 (41%)</td>
<td>27 (53%)</td>
</tr>
<tr>
<td>Infants tested at post-term age 12-13 weeks: n (%)</td>
<td>11 (22%)</td>
<td>13 (25%)</td>
<td>24 (47%)</td>
</tr>
</tbody>
</table>

SD: Standard deviation

Table 2. Intra-class correlation coefficient between Test 1 and Test 2 for all infants and for the two age groups when the tests were performed.

<table>
<thead>
<tr>
<th></th>
<th>ICC, 95% CI, Sw, SDD</th>
</tr>
</thead>
<tbody>
<tr>
<td>TIMPSI score for all infants (n=51)</td>
<td>0.99 (0.98-0.99)</td>
</tr>
<tr>
<td>TIMPSI score for infants tested at postmenstrual age 36-37 weeks (n=27)</td>
<td>0.94 (0.87-0.97)</td>
</tr>
<tr>
<td>TIMPSI score for infants tested at post-term age 12-13 weeks (n=24)</td>
<td>0.93 (0.84-0.97)</td>
</tr>
</tbody>
</table>

ICC: intra-class correlation coefficient, CI: confidence interval, Sw: measurement error, SDD: smallest detectable difference.
Figure 1. Bland Altman plot of the difference against the mean of the TIMPSI scores on Test 1 and Test 2 with mean difference in solid line and ± 1.96 SD (95% of agreement) in broken lines. Red circles represent infants from St. Olavs Hospital, Trondheim University Hospital, Norway, while blue circles represent infants from the University Hospital of North Norway.